Global Learning Methods for Latent Variable Sequence Models

Cem Subakan (PhD student)

University of Illinois at Urbana-Champaign

July 18'th, 2017

Outline

Introduction

Method of Moments

MoM Introduction Mixture of HMMs SHMM extension Bakis-HMM

Factorial HMM

Sequence Modeling

- E.g. Speech, Handwriting, Music, Text, Finance, and
- Uber



PANDARUS:

Alas, I think he shall be come approached and the day When little srain would be attain'd into being never fed, And who is but a chain and subjects of his death, I should not sleep.

Second Senator: They are away this miseries, produced upon my soul, Breaking and strongly should be buried, when I perish The earth and thoughts of many states.

Latent Variable Sequence Modeling

• HMM/LDS: $p(x_{1:T}) = \sum_{h_{1:T}} \prod_t p(x_t|h_t) p(h_t|h_{t-1})$

Latent Variable Sequence Modeling

• HMM/LDS: $p(x_{1:T}) = \sum_{h_{1:T}} \prod_{t} p(x_t|h_t) p(h_t|h_{t-1})$



Latent Variable Sequence Modeling

• HMM/LDS: $p(x_{1:T}) = \sum_{h_{1:T}} \prod_t p(x_t|h_t) p(h_t|h_{t-1})$



Fully Observed Sequence Modeling

Latent Variable Sequence Modeling

 $\blacktriangleright \text{HMM/LDS: } p(x_{1:T}) = \sum_{h_{1:T}} \prod_t p(x_t|h_t) p(h_t|h_{t-1})$



- Fully Observed Sequence Modeling
 - Markov Model $p(x_{1:T}) = \prod_t p(x_t | x_{t-1})$



Latent Variable Sequence Modeling

 $\blacktriangleright \text{HMM/LDS: } p(x_{1:T}) = \sum_{h_{1:T}} \prod_t p(x_t|h_t) p(h_t|h_{t-1})$



- Fully Observed Sequence Modeling
 - RNN $p(x_{1:T}) = \prod_t p(x_t | x_{1:t-1})$



Familiar Sequence Models

Mixture of Markov Models



[Subakan, et al. 2013] Mixture of HMMs





Maximum Likelihood is the first thing that comes to mind:

$$\max_{ heta} \log p(x| heta) = \max_{ heta} \log \sum_{h} p(x, h| heta)$$

Maximum Likelihood is the first thing that comes to mind:

$$\max_{ heta} \log p(x| heta) = \max_{ heta} \log \sum_{h} p(x, h| heta)$$

We can use Jensen's inequality by injecting a logarithm, and the distribution q(h):

$$\log \sum_{h} p(x, h|\theta) \frac{q(h)}{q(h)} = \log \mathbb{E}_{q(h)} \left[\frac{p(x, h|\theta)}{q(h)} \right]$$
$$\geq \mathbb{E}_{q(h)} \left[\log p(x, h|\theta) \right] + H_{q}$$

Maximum Likelihood is the first thing that comes to mind:

$$\max_{ heta} \log p(x| heta) = \max_{ heta} \log \sum_{h} p(x, h| heta)$$

We can use Jensen's inequality by injecting a logarithm, and the distribution q(h):

$$\log \sum_{h} p(x, h|\theta) \frac{q(h)}{q(h)} = \log \mathbb{E}_{q(h)} \left[\frac{p(x, h|\theta)}{q(h)} \right]$$
$$\geq \mathbb{E}_{q(h)} \left[\log p(x, h|\theta) \right] + H_q$$

This objective is in general not jointly convex.

Maximum Likelihood is the first thing that comes to mind:

$$\max_{ heta} \log p(x| heta) = \max_{ heta} \log \sum_{h} p(x, h| heta)$$

We can use Jensen's inequality by injecting a logarithm, and the distribution q(h):

$$egin{aligned} \log \sum_{h} p(x,h| heta) rac{q(h)}{q(h)} = \log \mathbb{E}_{q(h)} \left[rac{p(x,h| heta)}{q(h)}
ight] \ &\geq \mathbb{E}_{q(h)} \left[\log p(x,h| heta)
ight] + H_q \end{aligned}$$

- This objective is in general not jointly convex.
- Is there an alternative method which yields a global solution for this?

Maximum Likelihood is the first thing that comes to mind:

$$\max_{ heta} \log p(x| heta) = \max_{ heta} \log \sum_{h} p(x, h| heta)$$

We can use Jensen's inequality by injecting a logarithm, and the distribution q(h):

$$\log \sum_{h} p(x, h|\theta) \frac{q(h)}{q(h)} = \log \mathbb{E}_{q(h)} \left[\frac{p(x, h|\theta)}{q(h)} \right]$$
$$\geq \mathbb{E}_{q(h)} \left[\log p(x, h|\theta) \right] + H_q$$

- This objective is in general not jointly convex.
- Is there an alternative method which yields a global solution for this?
- ► (Probably) No. (P ≠ NP). But there are "close" problems which are easier to solve.





Disclaimer: We will not necessarily go to the red summit.

 Convex optimization problems. (Some convex problems are not poly-time solvable though)

- Convex optimization problems. (Some convex problems are not poly-time solvable though)
- Non-Convex canonical example:

$$\begin{split} \min_{U, \Sigma, V} \| X - U \Sigma V^\top \|_F \\ U^\top U &= I, \\ V^\top V &= I, \\ \Sigma &\geq 0, \ diagonal \end{split}$$

- Convex optimization problems. (Some convex problems are not poly-time solvable though)
- Non-Convex canonical example:

$$\begin{split} \min_{U, \Sigma, V} \| X - U \Sigma V^\top \|_F \\ U^\top U &= I, \\ V^\top V &= I, \\ \Sigma &\geq 0, \ diagonal \end{split}$$

Rayleigh Quotient

- Convex optimization problems. (Some convex problems are not poly-time solvable though)
- Non-Convex canonical example:

$$\begin{split} \min_{U, \Sigma, V} \| X - U \Sigma V^\top \|_F \\ U^\top U &= I, \\ V^\top V &= I, \\ \Sigma &\geq 0, \ diagonal \end{split}$$

- Rayleigh Quotient
- Procrustes Problem

- Convex optimization problems. (Some convex problems are not poly-time solvable though)
- Non-Convex canonical example:

$$\begin{split} \min_{U, \Sigma, V} \| X - U \Sigma V^\top \|_F \\ U^\top U &= I, \\ V^\top V &= I, \\ \Sigma &\geq 0, \ diagonal \end{split}$$

- Rayleigh Quotient
- Procrustes Problem
- Sinusoid Estimation (ESPRIT)

L

- Convex optimization problems. (Some convex problems are not poly-time solvable though)
- Non-Convex canonical example:

$$\begin{split} \min_{\substack{X \in V}} \| X - U \Sigma V^\top \|_F \\ U^\top U &= I, \\ V^\top V &= I, \\ \Sigma &\geq 0, \ diagonal \end{split}$$

- Rayleigh Quotient
- Procrustes Problem
- Sinusoid Estimation (ESPRIT)
- Method of Moments for Latent Variable Models

L

- Convex optimization problems. (Some convex problems are not poly-time solvable though)
- Non-Convex canonical example:

$$\begin{split} \min_{\substack{, \Sigma, V}} & \| X - U \Sigma V^\top \|_F \\ & U^\top U = I, \\ & V^\top V = I, \\ & \Sigma \ge 0, \text{ diagonal} \end{split}$$

- Rayleigh Quotient
- Procrustes Problem
- Sinusoid Estimation (ESPRIT)
- Method of Moments for Latent Variable Models
- Dictionary learning. (under assumptions)

Introduction

Method of Moments

MoM Introduction Mixture of HMMs SHMM extension Bakis-HMM

Factorial HMM

The other way: Method of Moments

The idea is to estimate the models parameters µ_{1:K} by solving a system of non-linear equations formed with moments E[g_k(x)], k ∈ {1,...K}:

$$\mathbb{E}[g_1(x)] = f_1(\mu_{1:K})$$
$$\vdots$$
$$\mathbb{E}[g_K(x)] = f_K(\mu_{1:K})$$

The other way: Method of Moments

The idea is to estimate the models parameters µ_{1:K} by solving a system of non-linear equations formed with moments E[g_k(x)], k ∈ {1,...K}:

$$\mathbb{E}[g_1(x)] = f_1(\mu_{1:\kappa})$$
$$\vdots$$
$$\mathbb{E}[g_{\kappa}(x)] = f_{\kappa}(\mu_{1:\kappa})$$

Canonical Example: x ~ G(a, b):

$$\begin{split} \mathbb{E}[x] &= ab & \longrightarrow & \widehat{b} = (\mathbb{E}[x^2] - \mathbb{E}[x]^2) / \mathbb{E}[x] \\ \mathbb{E}[x^2] &= ab^2 + a^2b^2 & \widehat{a} = \mathbb{E}[x]^2 / (\mathbb{E}[x^2] - \mathbb{E}[x]^2) \end{split}$$

The other way: Method of Moments

The idea is to estimate the models parameters µ_{1:K} by solving a system of non-linear equations formed with moments E[g_k(x)], k ∈ {1,...K}:

$$\mathbb{E}[g_1(x)] = f_1(\mu_{1:K})$$

$$\mathbb{E}[g_{\mathcal{K}}(x)] = f_{\mathcal{K}}(\mu_{1:\mathcal{K}})$$

Canonical Example: x ~ G(a, b):

$$\begin{split} \mathbb{E}[x] &= ab & \longrightarrow & \widehat{b} = (\mathbb{E}[x^2] - \mathbb{E}[x]^2) / \mathbb{E}[x] \\ \mathbb{E}[x^2] &= ab^2 + a^2b^2 & \widehat{a} = \mathbb{E}[x]^2 / (\mathbb{E}[x^2] - \mathbb{E}[x]^2) \end{split}$$

 This is not as statistically efficient as ML (CRLB). But the problem is (usually) "easier". MoM for spherical GMM: [Hsu, Kakade 13]



$$egin{array}{ll} h &\sim \ {\it Cat}(\pi) \ x | h &\sim \ {\cal N}(\mu_h, \sigma^2 I) \end{array}$$

Method of Moments for LVMs

MoM for spherical GMM: [Hsu, Kakade 13]



$$egin{array}{ll} h &\sim & {\it Cat}(\pi) \ x | h &\sim & {\cal N}(\mu_h, \sigma^2 I) \end{array}$$

Let's write down some moments:

$$\mathbb{E}[x] = \sum_{k=1}^{K} \mu_k \pi_k,$$

Method of Moments for LVMs

MoM for spherical GMM: [Hsu, Kakade 13]



$$egin{array}{ll} h &\sim & {\it Cat}(\pi) \ x | h &\sim & {\cal N}(\mu_h, \sigma^2 I) \end{array}$$

Let's write down some moments:

$$\mathbb{E}[x] = \sum_{k=1}^{K} \mu_k \pi_k, \ \mathbb{E}[x \otimes x] = \sum_{k=1}^{K} \pi_k \ \mu_k \otimes \mu_k + \sigma^2 I$$

Method of Moments for LVMs

MoM for spherical GMM: [Hsu, Kakade 13]



$$egin{array}{ll} h &\sim & {\it Cat}(\pi) \ x | h &\sim & {\cal N}(\mu_h, \sigma^2 I) \end{array}$$

Let's write down some moments:

$$\mathbb{E}[x] = \sum_{k=1}^{K} \mu_k \pi_k, \ \mathbb{E}[x \otimes x] = \sum_{k=1}^{K} \pi_k \ \mu_k \otimes \mu_k + \sigma^2 I$$
$$\mathbb{E}[x \otimes x \otimes x] = \sum_{k=1}^{K} \pi_k \ \mu_k \otimes \mu_k \otimes \mu_k$$
$$+ \sigma^2 \left(\sum_{l=1}^{L} \mathbb{E}[x] \otimes e_l \otimes e_l + e_l \otimes \mathbb{E}[x] \otimes e_l + e_l \otimes \mathbb{E}[x] \right)$$

MoM for spherical GMM: [Hsu, Kakade 13]

$$\mathbb{E}[x] = \sum_{k=1}^{K} \mu_k \pi_k, \ \mathbb{E}[x \otimes x] = \sum_{k=1}^{K} \pi_k \ \mu_k \otimes \mu_k + \text{garbage}$$
$$\mathbb{E}[x \otimes x \otimes x] = \sum_{k=1}^{K} \pi_k \ \mu_k \otimes \mu_k \otimes \mu_k + \text{garbage}$$

► Form the system of equations:

$$M_{2} := \mathbb{E}[x \otimes x] - \text{garbage} = \sum_{k=1}^{K} \pi_{k} \ \mu_{k} \otimes \mu_{k}$$
$$M_{3} := \mathbb{E}[x \otimes x \otimes x] - \text{garbage} = \sum_{k=1}^{K} \pi_{k} \ \mu_{k} \otimes \mu_{k} \otimes \mu_{k}$$

Form the system of equations:

$$M_{2} := \mathbb{E}[x \otimes x] - \text{garbage} = \sum_{k=1}^{K} \pi_{k} \ \mu_{k} \otimes \mu_{k}$$
$$M_{3} := \mathbb{E}[x \otimes x \otimes x] - \text{garbage} = \sum_{k=1}^{K} \pi_{k} \ \mu_{k} \otimes \mu_{k} \otimes \mu_{k}$$





Third Order Moment Tensor

Weighted sum of outer product of parameter vectors. • Whiten M_3 with a matrix W, such that:

 $W^{\top}M_2W = I$

• Whiten M_3 with a matrix W, such that:

$$W^{\top}M_2W = I$$

► Then the eigenvectors of

$$\widetilde{M}_3 = \sum_{k=1}^{K} w_k(W^{\top}\mu) \otimes (W^{\top}\mu) \otimes (W^{\top}\mu)$$

are obtainable via power iterations.
(and where do I come in)

- PCA
- ICA papers from 90s [mainly Cardoso]
- System ID literature from 90s. (Kalman Filters)
- ▶ Inference in HMMs [Hsu et al. 09]
- Parameter Estimation in HMMs [Anandkumar et al. 12, 14]
- Multiview Discrete/Mixture Models [Anandkumar et al. 12]
- ▶ Inference in general trees [Parikh et al. 12]
- Single View Spherical GMMs [Hsu, Kakade, 13]
- ▶ Parameter estimation in somewhat general graphs [Chaganty, Liang 14]
- Framework for HMMs with special transition structures. [Me et al., 14,15]
- Attempts on Neural Networks [Anandkumar, 15,16]

Spectral Learning of Mixture of HMMs

[Smyth, 97]



 $egin{aligned} h_n &\sim \textit{Categorical}(\pi_n)\ \mathbf{x}_n &\sim \textit{HMM}(A_n,O_n) \end{aligned}$

Spectral Learning of Mixture of HMMs

[Smyth, 97]



$$h_n \sim Categorical(\pi_n)$$

 $\mathbf{x}_n \sim HMM(A_n, O_n)$

• Learning Goal: Estimate π_n, A_n, O_n , given $\mathbf{x}_{1:N}$

Mixture of HMMs



 $egin{aligned} h_n &\sim \textit{Categorical}(\pi_n) \ \mathbf{x}_n | h_n &\sim \textit{HMM}(
ho_{h_n}, A_{h_n}, O_{h_n}) \end{aligned}$

Mixture of HMMs



Problem: The moment estimator is agnotic to the block structure of the model.

• An MHMM with *local* parameters $\theta_{1:K} = (O_{1:K}, A_{1:K}, \nu_{1:K}, \pi)$ is an HMM with *global* parameters $\bar{\theta} = (\bar{O}, \bar{A}, \bar{\nu})$, where:

$$\bar{O} = \begin{bmatrix} O_1 & \dots & O_K \end{bmatrix}, \quad \bar{A} = \begin{bmatrix} A_1 & \mathbf{0} & \dots & \mathbf{0} \\ \mathbf{0} & A_2 & \dots & \mathbf{0} \\ & & \ddots \\ \mathbf{0} & \mathbf{0} & \dots & A_K \end{bmatrix}, \quad \bar{\nu} = \begin{bmatrix} \pi_1 \nu_1 \\ \pi_2 \nu_2 \\ \vdots \\ \pi_K \nu_K \end{bmatrix}$$

How to impose this structural constraint on the estimator?

٠

HMM-Mixture model equivalence, [Kontorovich et al., 13]

An HMM with state marginals $p(h_t)$ is equivalent to a mixture model with mixing weights $\pi := \frac{1}{T} \sum_{t=1}^{T} p(h_t)$, and the same emission parameters.

HMM-Mixture model equivalence, [Kontorovich et al., 13]

An HMM with state marginals $p(h_t)$ is equivalent to a mixture model with mixing weights $\pi := \frac{1}{T} \sum_{t=1}^{T} p(h_t)$, and the same emission parameters.

- First compute (estimate) \widehat{O} , and \widehat{pi} .
- Then solve the convex problem:

$$\begin{split} \min_{A} & \|M_2 - \widehat{O}A \operatorname{diag}(\widehat{\pi}) \widehat{O}\|_F \\ s.t. & \mathbf{1}^\top A = \mathbf{1}^\top, \\ & A \geq 0. \end{split}$$

Two stage estimation framework

- Get rough/permuted estimates for the parameters $\widehat{O}, \widehat{A}, \widehat{\pi}$.
- De-permute A. (Solve the graph problem dictated by model)
- Solve:

$$\begin{split} \min_{A} & \|M_2 - \widehat{O}A \text{diag}(\widehat{\pi}) \widehat{O}\|_F \\ s.t. & 1^\top A = 1^\top, \\ & A \ge 0. \\ & f(\mathcal{M}, A) = 0 \end{split}$$

• f, and \mathcal{M} depend on the model.

The framework handles these models:

- ▶ **MHMM**: $f(\mathcal{M}, A) = A \odot (1 \mathcal{M}) = \mathbf{0}$. \mathcal{M} is block diagonal.
- ▶ SHMM: $f(\mathcal{M}, A) = A \odot (1 \mathcal{M}) \widehat{B} \otimes \frac{1}{M} \mathbf{1}_M \mathbf{1}_M^\top = \mathbf{0}$. \mathcal{M} is block diagonal.
- ▶ Left-to-Right HMM: f(M, A) = A ⊙ (1 − M) = 0, estimate M with a greedy graph traversal algorithm. M is lower triangular.
- Bakis HMM: f(M, A) = A ⊙ (1 − M) = 0, M corresponds to an Hamiltonian circuit (TSP approximation). M is binary lower first uni-triangular.
- HMM with mixture emissions: $f(\mathcal{M}, A^{i,j}) = A^{i,j} \mathbf{1}^{\top}$.

Mixture of HMMs: De-permutation

• $\lim_{e\to\infty} \bar{A}^e = [\bar{v}_1 1_M^\top, \bar{v}_2 1_M^\top, \ldots, \bar{v}_K 1_M^\top]$, where \bar{v}_k is the k'th eigenvector of \bar{A} .



Mixture of HMMs: De-permutation

• $\lim_{e\to\infty} \bar{A}^e = [\bar{v}_1 1_M^\top, \bar{v}_2 1_M^\top, \dots, \bar{v}_K 1_M^\top]$, where \bar{v}_k is the k'th eigenvector of \bar{A} .



• What happens for $\mathcal{P}(\bar{A})$:



Mixture of HMMs: De-permutation

• $\lim_{e\to\infty} \bar{A}^e = [\bar{v}_1 1_M^\top, \bar{v}_2 1_M^\top, \dots, \bar{v}_K 1_M^\top]$, where \bar{v}_k is the k'th eigenvector of \bar{A} .



• What happens for $\mathcal{P}(\bar{A})$:



What happens in practice:



MHMM De-permutation Continued





• Then form rank- \widehat{K} reconstruction A^r :

$$A^{r} = V_{1:\widehat{K}} \Lambda_{1:\widehat{K}} V^{-1}$$

Then Cluster. (A La Spectral Clustering)

Digit clustering with MHMMs:

Algorithm	1v2	1v3	1v4	1v5	2v3	2v4	2v5
Spectral	100	70	54	55	83	99	99
EM init. w/ Spectral	100	99	100	100	96	100	100
EM init. at Random	96	99	98	54	83	100	100



$$\begin{array}{rcl} h_t | h_{t-1} & \sim {\it Cat}(B(:,h_{t-1})) \\ r_t | r_{t-1},h_t,h_{t-1} & \sim [h_t = h_{t-1}] {\it Cat}(A(:,r_{t-1},h_t)) \\ & + [h_t \neq h_{t-1}] {\it U}(.) \\ x_t | h_t,r_t & \sim p(x_t | h_t,r_t) \end{array}$$

$$\bar{O} = \begin{bmatrix} O_1 & \dots & O_K \end{bmatrix}, \bar{A} = \begin{bmatrix} B_{1,1}A_1 & B_{1,2}\frac{11^\top}{M} & \dots & B_{1,M}\frac{11^\top}{M_1} \\ B_{2,1}\frac{11^\top}{M} & B_{2,2}A_2 & \dots & B_{2,M}\frac{11^\top}{M} \\ & \ddots & \\ B_{M,1}\frac{11^\top}{M} & B_{M,2}\frac{11^\top}{M} & \dots & B_{M,M}A_K \end{bmatrix},$$
$$\bar{\nu} = \begin{bmatrix} \pi_1\nu_1 & \pi_2\nu_2 & \dots & \pi_K\nu_K \end{bmatrix}^\top.$$

How to impose this structural constraint on the estimator?

$$\bar{O} = \begin{bmatrix} O_1 & \dots & O_K \end{bmatrix}, \bar{A} = \begin{bmatrix} B_{1,1}A_1 & B_{1,2}\frac{11^{\top}}{M} & \dots & B_{1,M}\frac{11^{\top}}{M} \\ B_{2,1}\frac{11^{\top}}{M} & B_{2,2}A_2 & \dots & B_{2,M}\frac{11^{\top}}{M} \\ & \ddots & \\ B_{M,1}\frac{11^{\top}}{M} & B_{M,2}\frac{11^{\top}}{M} & \dots & B_{M,M}A_K \end{bmatrix},$$
$$\bar{\nu} = \begin{bmatrix} \pi_1\nu_1 & \pi_2\nu_2 & \dots & \pi_K\nu_K \end{bmatrix}^{\top}.$$

- How to impose this structural constraint on the estimator?
- Use the same de-permutation method as MHMM.

MHMM-SHMM spectrum

•
$$A_{i,i} \sim \text{Dirichlet}(1,\ldots,1), B = \begin{bmatrix} \alpha & 1-\alpha \\ 1-\alpha & \alpha \end{bmatrix}.$$



Bakis-HMM

▶ Is an HMM that can only move one state at a time.



- Every state is visited exactly once.
- Depermutation: Find a maximum weight Hamiltonian circuit on Â. (Traveling Salesman problem)



Experimental Results



Impressions on MoM:

► Good:

- Global
- ▶ Initialization: No need to worry about initialization. Also can initialize EM.
- Scalable: Computationally cheap: Gather the moments, factorize a small matrix.
- Interesting/Theoretical: Bounds.
- **Subroutine:** Potentially can be used as a subroutine under EM.

Impressions on MoM:

► Good:

- Global
- ▶ Initialization: No need to worry about initialization. Also can initialize EM.
- Scalable: Computationally cheap: Gather the moments, factorize a small matrix.
- Interesting/Theoretical: Bounds.
- **Subroutine:** Potentially can be used as a subroutine under EM.
- ► Bad:
 - Model Mismatch: Horrible in regards to model mismatch. (Hard assumption on model Unlike ML, which minimizes KL(p||q).
 - Not as statistically efficient as ML.

Impressions on MoM:

► Good:

- Global
- ▶ Initialization: No need to worry about initialization. Also can initialize EM.
- Scalable: Computationally cheap: Gather the moments, factorize a small matrix.
- Interesting/Theoretical: Bounds.
- **Subroutine:** Potentially can be used as a subroutine under EM.
- ► Bad:
 - Model Mismatch: Horrible in regards to model mismatch. (Hard assumption on model Unlike ML, which minimizes KL(p||q).
 - Not as statistically efficient as ML.
- ► Ugly:
 - You can get complex numbers for parameter estimates/likelihoods.

Introduction

Method of Moments

MoM Introduction Mixture of HMMs SHMM extension Bakis-HMM

Factorial HMM

Factorial HMM

[Ghahramani, Jordan; 97]



Factorial HMM

[Ghahramani, Jordan; 97]





The dictionary Activations noise

Some Dictionary Learning Perspective..

General Dictionary Learning



- PCA: Both O and R are orthogonal.
- ▶ ICA: Solvable if *R* has independent coordinates.
- Mixture Model: *R* is one sparse. Solvable is *O* has full column rank.
- ▶ Sparse Dictionary Learning: Solvable if *O* is square and *R* is sparse Bernouilli-Gaussian. [Spielman et al. 12]

Some Dictionary Learning Perspective..

General Dictionary Learning



- PCA: Both O and R are orthogonal.
- ▶ ICA: Solvable if *R* has independent coordinates.
- Mixture Model: *R* is one sparse. Solvable is *O* has full column rank.
- ▶ Sparse Dictionary Learning: Solvable if *O* is square and *R* is sparse Bernouilli-Gaussian. [Spielman et al. 12]

Factorial Models:

$$O = \begin{bmatrix} O^1 & \dots & O^K \end{bmatrix}, \ R = \begin{bmatrix} R^1 \\ \vdots \\ R^K \end{bmatrix}$$

- ▶ No constraint on *O*, columns of *R* are block-*K* sparse.
- No Unique Solution!!!

Rank Deficiency

 $\operatorname{rank}(R) \leq MK - (K - 1)$

Rank Deficiency

 $\mathsf{rank}(R) \leq MK - (K - 1)$

Proof Sketch:

 $\dim(\operatorname{null}(R^{\top})) \geq K - 1.$

Therefore from rank-nullity theorem rank(R) = MK - (K - 1).

FHMM is unidentifiable

For a given assignment matrix $R \in \mathbb{R}^{KM \times T}$ There exists $O_1 \neq O_2$ such that $\prod_t \mathcal{N}(x_t | O_1 R, \sigma^2 I) = \prod_t \mathcal{N}(x_t | O_2 R, \sigma^2 I)$.

Rank Deficiency

 $\mathsf{rank}(R) \leq MK - (K - 1)$

Proof Sketch:

 $\dim(\operatorname{null}(R^{\top})) \geq K - 1.$

Therefore from rank-nullity theorem rank(R) = MK - (K - 1).

FHMM is unidentifiable

For a given assignment matrix $R \in \mathbb{R}^{KM \times T}$ There exists $O_1 \neq O_2$ such that $\prod_t \mathcal{N}(x_t | O_1 R, \sigma^2 I) = \prod_t \mathcal{N}(x_t | O_2 R, \sigma^2 I)$.

Proof: Since dim(null(R^{\top})) $\geq K - 1$, $(O_1 - O_2)R = 0$, for $O_1 \neq O_2$.

Shared Component FM

$$\forall k, \ O^{k} = \begin{bmatrix} | & | & | & | & | \\ \mu_{k}^{1} & \mu_{2}^{k} & \dots & \mu_{M-1}^{k} & s \\ | & | & | & | & | \end{bmatrix}$$

SC-FM is identifiable

Given an assignment matrix \tilde{R} which is rank MK - (K - 1), the emission matrix of an SC-FM is identifiable.

Proof Sketch:

 $\dim(\operatorname{null}(R^{\top})) = 0.$

Therefore $(O_1 - O_2)R \neq 0$, $\forall O_1 \neq O_2$.

Learning Example for Shared Component Factorial Model

• **Gist:** If the shared component *s* is incoherent, then we can identify it, and reveal the other components.

Example Observations



Obtained Components with SC-FM

5201

Components with regular model-EM



Learning Example for Shared Component Factorial Model

• **Gist:** If the shared component *s* is incoherent, then we can identify it, and reveal the other components.

Example Observations



Obtained Components with SC-FM

5201 6341

Components with regular model-EM



The shared component + incoherence assumption a bit too restrictive. Can we think of another model?

FHMM Identifiable Alternative 2



Identifiability follows similarly from the activation matrix R.
Revealing FHMM Practical Algorithm

Practical Algorithm for Revealing FHMM

- Cluster the data matrix $X \in \mathbb{R}^{L \times T}$ into clusters $X^c \in \mathbb{R}^{L \times C}$.
- Solve:

$$\begin{split} \min_{H} & \|X^{c} - X^{c}H\|_{F}^{2} + \beta \|H\|_{1}, \\ s.t. \ H_{i,i} &= 0, \text{ for } 1 \leq i \leq C, \\ & H \geq 0, \end{split}$$

where $H \in \mathbb{R}^{C \times C}$.

• Construct a bi-partite graph by reading the solution for *H*.

Revealing FHMM Practical Algorithm

Practical Algorithm for Revealing FHMM

- Cluster the data matrix $X \in \mathbb{R}^{L \times T}$ into clusters $X^c \in \mathbb{R}^{L \times C}$.
- Solve:

$$\begin{split} \min_{H} & \|X^{c} - X^{c}H\|_{F}^{2} + \beta \|H\|_{1}, \\ s.t. \ H_{i,i} &= 0, \text{ for } 1 \leq i \leq C, \\ & H \geq 0, \end{split}$$

where $H \in \mathbb{R}^{C \times C}$.

- Construct a bi-partite graph by reading the solution for H.
- ▶ Condition for learnability: Let O₁ = [x₀, x₁], O₂ = [y₀, y₁]. Observed combinations needs to form a connected bi-partite graph (Connectivity) (*linear number of edges in number of components, not quadratic*), and we need to observe all nodes and edges (Observability).



Unsupervised audio source separation example

- ▶ We mixed recording of double bass and flute (at 0dB).
- The observed mixtures satisfy the connectivity constraint.

Original Mixture



Reconstruction

We obtain almost perfect source separation.



▶ The algorithm is robust to the choice of number of clusters *C*.

- Identifiability: The original FHMM model is unidentifiable.
- ► Identifiable Alternatives: There exists identifiable alternatives which are globally learnable under stringent assumptions.
- ► Unsupervised Source Separation: Revealing FHMM works well under the connectedness and observability assumptions.
- Future work:
 - Can we relax the observability assumption so that we only require to observe less nodes in the connectivity graph?
 - Potential application in semi-supervised source separation.