#### The Qualification Exam

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UIUC CS

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# Outline



# My background

My research

- Paper 1: Estimating Latent Variable Graphical Models using Moments and Likelihoods
  - Introduction
  - Intro to method of moments for LVMs
  - The paper
- Second Paper, The Visual Microphone: Passive Recovery of Sound from Video
  - Introduction, The problem setup
  - Processing step

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#### • My research in one sentence:

I like big algorithms for small data, and I like NIPS/ICML style machine learning.

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- In particular, I am working on Method of Moments (MoM) for parameter estimation in LVMs. (Also known as Spectral Learning). Score so far:
  - M.Sc. Thesis
  - 2 NIPS workshop papers
  - 1 journal paper
  - NIPS 2014 paper NEW!

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- WHY MoM estimators?
  - They are cool, mathy and new (hip).
  - Avoid the everlasting local optima issue. (No initialization!)
  - Computationally much more efficient.
  - Learning guarantees.

#### Recent Research

- M.Sc. Thesis: Two new MoM algorithms for time series clustering.
- ICML 2014 submission: A Non-Negative Matrix Factorization (NMF) based framework for learning HMM variants with MoM.



Switching HMM

Factorial HMM

• Paper accepted to NIPS 2014! (acceptance rate: 414/1678) A Method of moments algorithm to learn mixture of HMMs.



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- It is unclear how to use standard MoM algorithms for this model.
- However, we can learn an HMM with MoM.

- Key idea: Mixture of HMMs is an HMM with block diagonal transition matrix.

$$\bar{O} = \begin{bmatrix} O_1 & \dots & O_K \end{bmatrix}, \quad \bar{A} = \begin{bmatrix} A_1 & \mathbf{0} & \dots & \mathbf{0} \\ \mathbf{0} & A_2 & \dots & \mathbf{0} \\ & \ddots & \\ \mathbf{0} & \mathbf{0} & \dots & A_K \end{bmatrix}, \quad \bar{\nu} = \begin{bmatrix} \pi_1 \nu_1 \\ \pi_2 \nu_2 \\ \vdots \\ \pi_K \nu_K \end{bmatrix}$$

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- The problem: Arbitrary permutation on parameter estimates, Parameters of different clusters get mixed up.
- Remedy: Block diagonal structure / spectral properties of the "global" transition matrix.

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• In real world, we have noise on off-block diagonal elements. This results in a global stationary distribution.



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Can we recover a block diagonal structure despite the estimation noise?



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- If the noise is not too severe, then yes we can. (Experimental and theoretical justification)
- Notice: Given the moments, computational burden does not depend on dataset size! (Unlike EM) COOL

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- Standard MoM algorithms are not directly applicable to models beyond HMM, GMM, LDA.
- This work proposes a framework for learning general graphical models.
- They divide the problem into two (three) stages, which helps generalizing.

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#### Problem Definition

• Let's suppose we have the following graphical model:



$$egin{aligned} &h\sim \mathsf{Discrete}(\pi_{1:K})\ &x_1|h\sim \mathcal{N}(\mu_{1,h}, \Sigma_1)\ &x_2|h\sim \mathcal{N}(\mu_{2,h}, \Sigma_2)\ &x_3|h\sim \mathcal{N}(\mu_{3,h}, \Sigma_3) \end{aligned}$$

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• Given  $\{x_{1,n}, x_{2,n}, x_{3,n}\}_{n=1}^N$ , can we estimate  $\mu_{1,1:K}, \mu_{2,1:K}, \mu_{3,1:K}$ ?

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• Given  $\{x_{1,n}, x_{2,n}, x_{3,n}\}_{n=1}^{N}$ , can we estimate  $\mu_{1,1:K}, \mu_{2,1:K}, \mu_{3,1:K}$ ? Yes we can!

#### The conventional way: EM

• Maximum Likelihood is the first thing that comes to mind:

$$\max_{\mu_{1:3}} p(x_{1:3,1:N}|\mu_{1:3}) = \max_{\mu_{1:3}} \sum_{h_{1:N}} p(x_{1:3,1:N}, h_{1:N}|\mu_{1:3})$$

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We can use Jensen's inequality by injecting a logarithm, and the distribution q(h<sub>1:N</sub>):

$$\log \sum_{h_{1:N}} p(x_{1:3,1:N}, h_{1:N} | \mu_{1:3}) \frac{q(h_{1:N})}{q(h_{1:N})} = \log \mathbb{E}_{q(h_{1:N})} \left[ \frac{p(x_{1:3,1:N}, h_{1:N} | \mu_{1:3})}{q(h_{1:N})} \right]$$
$$\geq \mathbb{E}_{q(h_{1:N})} \left[ \log p(x_{1:3,1:N}, h_{1:N} | \mu_{1:3}) \right] + H_q$$

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q(h<sub>1:N</sub>) = p(h<sub>1:N</sub>|x<sub>1:N</sub>, μ<sub>1:3</sub>) in EM. In E step q is updated. In M step we maximize this lower bound. It is obviously prone to local optima.

#### The other way: Method of Moments

The idea is to estimate the models parameters µ<sub>1:K</sub> by solving a system of non-linear equations formed with moments E[g<sub>k</sub>(x)], k ∈ {1,...K}:

$$\mathbb{E}[g_1(x)] = f_1(\mu_{1:K})$$
$$\vdots$$
$$\mathbb{E}[g_K(x)] = f_K(\mu_{1:K})$$
The idea is to estimate the models parameters µ<sub>1:K</sub> by solving a system of non-linear equations formed with moments 𝔼[𝑔<sub>k</sub>(𝑥)], k ∈ {1,...K}:

$$\mathbb{E}[g_1(x)] = f_1(\mu_{1:\kappa})$$
$$\vdots$$
$$\mathbb{E}[g_{\kappa}(x)] = f_{\kappa}(\mu_{1:\kappa})$$

• Canonical Example:  $x \sim \mathcal{G}(a, b)$ :

 $\mathbb{E}[x] = ab \qquad \rightarrow \qquad \widehat{b} = (\mathbb{E}[x^2] - \mathbb{E}[x]^2) / \mathbb{E}[x]$  $\mathbb{E}[x^2] = ab^2 + a^2b^2 \qquad \qquad \widehat{a} = \mathbb{E}[x]^2 / (\mathbb{E}[x^2] - \mathbb{E}[x]^2)$ 



 $egin{aligned} &h_n \sim \mathsf{Discrete}(\pi) \ &x_1 | h \sim \mathcal{N}(\mu_{1,h}, \Sigma_1) \ &x_2 | h \sim \mathcal{N}(\mu_{2,h}, \Sigma_2) \ &x_3 | h \sim \mathcal{N}(\mu_{3,h}, \Sigma_3) \end{aligned}$ 

• Let's write down some moments:

$$P_2 := \mathbb{E}[x_1 \otimes x_2] = \sum_{h=1}^K \pi_h \mathbb{E}[x_1|h] \otimes \mathbb{E}[x_2|h] = \sum_{h=1}^K \pi_h \ \mu_{1,h} \otimes \mu_{2,h}$$



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• So,  $P_2 = M_1 \text{diag}(\pi) M_2$  and  $P_{3,i} = M_1 \text{diag}(M_3(i,:)) \text{diag}(\pi) M_2$ .



$$\begin{split} h_n &\sim \mathsf{Discrete}(\pi) \\ x_1 | h &\sim \mathcal{N}(\mu_{1,h}, \Sigma_1) \\ x_2 | h &\sim \mathcal{N}(\mu_{2,h}, \Sigma_2) \\ x_3 | h &\sim \mathcal{N}(\mu_{3,h}, \Sigma_3) \end{split}$$

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• So,  $P_2 = M_1 \operatorname{diag}(\pi)M_2$  and  $P_{3,i} = M_1 \operatorname{diag}(M_3(i, .)) \operatorname{diag}(\pi)M_2$ • And,  $P_{3,i}P_2^{-1} = M_1 \operatorname{diag}(M_3(i, .))M_1^{-1}$ , which is an eigenvalue decomposition (assuming invertibility). < □ > < 同 17 / 40

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- And, P<sub>3,i</sub>P<sub>2</sub><sup>-1</sup> = M<sub>1</sub>diag(M<sub>3</sub>(i,:))M<sub>1</sub><sup>-1</sup>, which is an eigenvalue decomposition (assuming invertibility).
- This is from Anandkumar et al. 2012, COLT paper. There are statistically more efficient ways now. (Using all three slices instead of one. Anandkumar et al. 2014, to appear in JMLR)

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• Now, can we learn A,  $\mu_1, \mu_2, \mu_3, \mu_4$  using moments?

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- Now, can we learn A,  $\mu_1, \mu_2, \mu_3, \mu_4$  using moments?
- Not straightforwardly with original work. But this paper says,

Yes, we can!

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# Paper 1: Key Idea

• Learning conditional moments and hidden marginals separately



• The pipeline:

- ► First estimate the conditional moments 𝔼[x<sub>i</sub>|h<sub>k</sub>].
- Then obtaining the hidden potential p(h<sub>2</sub>|h<sub>1</sub>) is easy.



3. 3

#### Part 1: Estimating the conditional moments



 Notice, h<sub>1</sub> has three conditionally independent "views". Thus, we can estimate E[x<sub>1</sub>|h<sub>1</sub>], E[x<sub>2</sub>|h<sub>1</sub>] and E[x<sub>3</sub>|h<sub>1</sub>].

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- $h_2$  has  $x_2, x_3, x_4$ . So,  $\mathbb{E}[x_2|h_2]$ ,  $\mathbb{E}[x_3|h_2]$  and  $\mathbb{E}[x_4|h_2]$  are available.

#### Part 1: Estimating the conditional moments



- Notice, h₁ has three conditionally independent "views". Thus, we can estimate E[x₁|h₁], E[x₂|h₁] and E[x₃|h₁].
- $h_2$  has  $x_2, x_3, x_4$ . So,  $\mathbb{E}[x_2|h_2]$ ,  $\mathbb{E}[x_3|h_2]$  and  $\mathbb{E}[x_4|h_2]$  are available.

$$\mathbb{E}[x_1 \otimes x_2 \otimes x_3] = \sum_{h_1} \sum_{h_2} p(h_1) p(h_2 | h_1) \mathbb{E}[x_1 | h_1] \mathbb{E}[x_2 | h_1] \mathbb{E}[x_3 | h_2]$$
  
=  $\sum_{h_1} p(h_1) \mathbb{E}[x_1 | h_1] \mathbb{E}[x_2 | h_1] \left( \sum_{h_2} p(h_2 | h_1) \mathbb{E}[x_3 | h_2, h_1] \right)$   
=  $\sum_{h_1} p(h_1) \mathbb{E}[x_1 | h_1] \mathbb{E}[x_2 | h_1] \mathbb{E}[x_3 | h_1] \rightarrow \text{Right form for MoM}$ 

#### Part 2: Estimating the hidden potentials



• Given  $\mathbb{E}[x_2|h_1]$  and  $\mathbb{E}[x_3|h_2]$ , estimating  $p(h_2, h_1)$  is child's play.

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For example,

$$\mathbb{E}[x_2 \otimes x_3] = \sum_{h_1, h_2} \mathbb{E}[x_2|h_1]p(h_2, h_1)\mathbb{E}[x_3|h_2]$$
$$= M_2 S M_3$$

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$$= M_2 S M_3$$

• One way to do it is convex optimization.

• If we choose to,  $\min_{S} \|\mathbb{E}[x_2 \otimes x_3] - M_2 S M_3\|_F$ , then the solution is  $\widehat{S} = M_2^{\dagger} \mathbb{E}[x_2 \otimes x_3] M_3^{\dagger}$ . (This is the first thing they do in the paper)

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- Or better we can do,

$$\begin{split} \min_{S} \left\| \mathbb{E}[x_2 \otimes x_3] - M_2 S M_3 \right\|_F \\ S \geq 0 \\ 1^T S 1 = 1 \end{split}$$

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(I use CVX! haha!)

- If we choose to,  $\min_{S} \|\mathbb{E}[x_2 \otimes x_3] M_2 S M_3\|_F$ , then the solution is  $\widehat{S} = M_2^{\dagger} \mathbb{E}[x_2 \otimes x_3] M_3^{\dagger}$ . (This is the first thing they do in the paper)
- Or better we can do,

$$\begin{split} \min_{S} \left\| \mathbb{E}[x_{2} \otimes x_{3}] - M_{2}SM_{3} \right\|_{F} \\ S \geq 0 \\ 1^{T}S1 = 1 \end{split}$$

(I use CVX! haha!)

• Or even better (so they claim),

$$\max_{S} \mathbb{E}[\log p(x_2, x_3)]$$

This is called the "Composite Likelihood"

# A simulation for computational and statistical efficiency

• Statistical and computational efficiencies of the two stage estimation and EM for HMM with Gaussian observations. (K = 5)



Conditions for recoverability of a Directed Graphical Model

- We need to be able to recover all conditional expectations:
  - Every hidden node must be a "bottleneck" in the worst case.
  - There must be at least three cond. indep. variables for a node to be a bottleneck.
  - > The conditional expectation matrices have to have full column rank.

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- > The conditional expectation matrices have to have full column rank.
- Examples:



Conditions for recoverability of a Directed Graphical Model

- Hidden nodes must possess the "Exclusive Views" property.
  - A hidden node has to have at least one conditionally independent observation on its own to have this property.
- If we want to estimate all hidden potentials:



PASS



# Part 3: Undirected Graphs (MRFs)

• The joint distribution is defined with clique "potentials".

$$p(h_{1:K}, x_{1:J}|\theta) = \frac{1}{Z(\theta)} \prod_{C \in \mathcal{G}} \exp(\theta^T \phi(x_C, h_C))$$

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• Example: (An image segmentation model)



$$Z(\theta) = \int \prod_{C \in \mathcal{G}} \exp(\theta^T \phi(x_C, h_C)) dx_{1:J} dh_{1:K}$$

The notorious partition function!

• The lower bound on likelihood is:

 $\log p(x_{1:K}|\theta) \geq \mathbb{E}_{p(x_{1:J},h_{1:K}|\theta)}[\log p(x_{1:J},h_{1:K}|\theta)] = \mathcal{L}(\theta)$ 

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- With MoM, we can estimate  $p(x_{1:J}, h_{1:K}|\theta)$  from data.

$$\mathcal{L}(\theta) = \theta^{T} \left( \sum_{C \in \mathcal{G}} \mathbb{E}[\phi(x_{1:J}, h_{1:K})] \right) - \mathcal{A}(\theta)$$

where, 
$$\mathbb{E}[\phi(x_{1:J}, h_{1:K})] = \sum_{x_{1:J}, h_{1:K}} p(x_{1:J}, h_{1:K}) \phi(x_{1:J}, h_{1:K})$$

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• So, the MoM lower bound is concave w.r.t.  $\theta$ .

#### Conclusions

- It's a good paper, that opens new possibilities for MoM learning.
- The moral of the story: MoM and likelihood maximization can be used synergistically to learn a variety of models.
- The story isn't finished yet: Models like MHMM is not covered. (where not all variables are bottlenecks.)
- Experimental verification is necessary as follow-up work.

# Outline

#### **1** Me

- My background
- My research
- Paper 1: Estimating Latent Variable Graphical Models using Moments and Likelihoods
  - Introduction
  - Intro to method of moments for LVMs
  - The paper

# 3 Second Paper, The Visual Microphone: Passive Recovery of Sound from Video

- Introduction, The problem setup
- Processing step

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 We'll mostly be interested in "Processing" step, which is somewhat involved in signal processing/vision.

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  - It is a filter bank consisting of sombrero type of filters with different orientations and scales.
- A similar filter bank, Gabor Wavelets (real parts) for several scales (r): and orientations θ:



• In 1D, it's of form  $f(x; \sigma^2, \omega) = \mathcal{N}(x, 0, \sigma^2) e^{j2\pi\omega x}$ 

• After wavelet transform, we have:

$$W(V) = \underbrace{A(r, \theta, x, y, t)}_{amplitude} e^{j \overline{\psi(r, \theta, x, y, t)}}$$

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- This is a phasor representation, A(.) is the amplitude and  $\psi(.)$  is the phase.
- Then phase variations wrt. to a reference frame  $t_0$  is computed  $\psi_v(., t) = \psi(., t) \psi(., t_0)$ .
  - For small motions these variations

This is the local motion signal.

- The output of this stage is the reconstruction!.
- First average over the spatial coordinates:

$$\Phi(r,\theta,t) = \sum_{x,y} A(r,\theta,x,y,t)^2 \psi_v(r,\theta,x,y,t)$$

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$$\Phi(r,\theta,t) = \sum_{x,y} A(r,\theta,x,y,t)^2 \psi_v(r,\theta,x,y,t)$$

• Then align the signals:

$$t_i = \arg \max_{t_i} \Phi(r_0, \theta_0, t)^T \Phi(r_i, \theta_i, t - t_i)$$

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• The reconstructed signal is:

$$\widehat{s}(t) = \sum_{i} \Phi(r_i, \theta_i, t - t_i)$$

Say we have the following video..

(Loading Video...)

#### • Can we reconstruct a sound?

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## Yes we can!



- Original signal is  $x(t) = \sin(2\pi\omega_0 t) + \sin(2\pi 2\omega_0 t)$
- I corrupt the original signal in each dimension with  $a x(t + \theta) + \epsilon$ ,  $a \sim \mathcal{U}([0 \ 1])$ ,  $\theta, \epsilon \sim \mathcal{N}(0, 0, 1)$



# A glance at their experiments

- Objects behave like low-pass filters. It's harder to obtain high frequencies, as one would expect.
- For speech, their method generally works worse than an active method.
- They claim that unintelligible sound may also be useful for surveillance type applications.
- They have the vibration mode estimation application also.
- Limitations: Sampling rate / Magnification

#### Conclusions

- A (very) good paper with lots of experiments.
- I would have liked to see some theoretical justification for the processing step.
- Experiments are really good, and they provide several applications, and some analysis. It's definitely a well studied, exciting (even for me) paper.